

# 1           **Global COVID-19 transmission rate is influenced by** 2           **precipitation seasonality and the speed of climate** 3           **temperature warming**

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9

## 10   **Abstract**

11   The novel coronavirus disease 2019 (COVID-19) became a rapidly spreading worldwide  
12   epidemic; thus, it is a global priority to reduce the speed of the epidemic spreading.  
13   Several studies predicted that high temperature and humidity could reduce COVID-19  
14   transmission. However, exceptions exist to this observation, further thorough  
15   examinations are thus needed for their confirmation. In this study, therefore, we used a  
16   global dataset of COVID-19 cases and global climate databases and comprehensively  
17   investigated how climate parameters could contribute to the growth rate of COVID-19  
18   cases while statistically controlling for potential confounding effects using spatial  
19   analysis. We also confirmed that the growth rate decreased with the temperature;  
20   however, the growth rate was affected by precipitation seasonality and warming velocity  
21   rather than temperature. In particular, a lower growth rate was observed for a higher  
22   precipitation seasonality and lower warming velocity. These effects were independent of  
23   population density, human life quality, and travel restrictions. The results indicate that the  
24   temperature effect is less important compared to these intrinsic climate characteristics,  
25   which might thus be useful for explaining the exceptions. However, the contributions of  
26   the climate parameters to the growth rate were moderate; rather, the contribution of travel  
27   restrictions in each country was more significant. Although our findings are preliminary  
28   owing to data-analysis limitations, they may be helpful when predicting COVID-19  
29   transmission.

## 30 **1. Introduction**

31 The world-wide spreading of coronavirus disease 2019 (COVID-19) [1], an infectious  
32 disease caused by the novel coronavirus, severe acute respiratory syndrome coronavirus 2  
33 (SARS-CoV-2 / 2019-nCoV) was firstly identified in Wuhan, China [2]. The COVID-19  
34 epidemic has a serious impact on the public health and economy [3], the reduction of its  
35 spreading is thus a significant challenge. How climate parameters are associated with the  
36 spreading is intriguing concerning the coronavirus characterization and spreading  
37 prediction. Previous studies have suggested that temperature increase could reduce  
38 COVID-19 transmission both in China [4–6] and at the global scale [7–11]. However, a  
39 bell-shaped or quadratic relationship between the COVID-19 transmission rate and the  
40 temperature was observed, indicating that the optimal transmission temperature could be  
41 at  $\sim 8$  °C. Moreover, part of the previous studies [4–6, 8] also reported that higher  
42 humidity is also associated with a lower transmission rate of COVID-19. These results  
43 are consistent with the influenza seasonality (i.e., the fact that influenza transmission is  
44 reduced due to temperature and humidity increase) [12]. Thus, previous studies have  
45 predicted that the arrival of summer and the rainy season would reduce COVID-19  
46 transmission.

47 However, more careful examinations are required to conclude such COVID-19  
48 seasonality. As emphasized in part of the previous studies, temperature could account for  
49 a relatively modest amount of the total variation in the COVID-19 transmission rate [10].  
50 In fact, despite the expectations, the spreading of COVID-19 has also been observed in  
51 warm and humid areas (e.g., Australia, Brazil, and Argentina, on Southern Hemisphere, in  
52 early March). This indicates that other climate parameters might also affect COVID-19  
53 transmission. For example, influenza transmission is also influenced by several  
54 environmental parameters, such as ultraviolet (UV) radiation, wind speed, precipitation,  
55 and air pollution. [13]; moreover, it also correlated with diurnal temperature ranges [14]  
56 and urbanization (human impacts) [15]. In addition to this, changing rapid weather  
57 variability (e.g., climate seasonality and climate change) increases the risk of an influenza  
58 epidemic [16]. In general, seasonal variations in temperature, rainfall, and resource  
59 availability can exert strong pressure on infectious disease population dynamics [17].  
60 Inspired by these results, previous studies evaluated the contributions of wind speed [8],  
61 precipitation and UV irradiation to COVID-19 transmission [9]. However, the remaining  
62 parameters have still been poorly investigated to date. In particular, the temperature  
63 might be associated with other climate parameters, it is thus necessary to control the  
64 potentially confounding effects [18–20].

65 To study the aforementioned subject, the application of spatial analysis might also be  
66 needed. Although spatial autocorrelations between observation areas and variables need  
67 to be evaluated when analyzing geographic data [18, 20, 21], previous studies have  
68 understudied them. It remains possible that the observed associations of COVID-19  
69 transmission with temperature and humidity are spatial autocorrelation artefacts.

70 In this study, we thus aimed at conducting a more comprehensive investigation. Using  
71 global time-series data on confirmed COVID-19 cases [1] and global climate databases,  
72 we comprehensively investigated how climate parameters contribute to COVID-19  
73 transmission on a global scale while statistically controlling for potential confounding

74 effects using spatial analysis. Population density and quality of human life (human  
75 development index) were also considered when controlling for potential confounding  
76 effects because they might affect infectious disease transmission [15], including COVID-  
77 19 transmission [4]. Similarly, we also considered the travel restrictions because the  
78 national emergency response, including travel bans, appears to have delayed the growth  
79 and limited the size of the COVID-19 epidemic in China [22, 23].

## 80 **2. Material and methods**

### 81 *2.1. The growth rate of COVID-19 cases*

82 We obtained global time-series data for the period between January 22, 2020 - April 6,  
83 2020, on the number of confirmed cases of COVID-19 [1] operated by the Johns Hopkins  
84 University Center for Systems Science and Engineering from their GitHub repository. In  
85 this repository, the global dataset and dataset of the (USA) were available. We combined  
86 these datasets after removing USA-related data from the global dataset. To estimate the  
87 COVID-19 transmission rate, many previous studies considered the measures based on  
88 the number of cases. However, it remains possible that the differences in the number of  
89 tested individuals between areas (countries) affect these measures. We thus used instead  
90 the growth rate of confirmed COVID-19 cases as a more suitable measure. The growth  
91 rate in each observation was computed using the R statistical software (version 3.6.2;  
92 [www.r-project.org](http://www.r-project.org)) and the package *incidence* (version 1.7.1) [24]; in particular, the *fit*  
93 function was used. To estimate the growth rate during the initial (exponential) phase, we  
94 used the data within 15 days (~2 weeks) starting from the date (call *first date* hereafter)  
95 when 30 and more cases were confirmed in cumulative counts, as described previously  
96 [7]. We confirmed that similar conclusions were obtained at the different cut-off values  
97 (using the data within 30 days starting from the date when 50 and more cases were  
98 confirmed).

### 99 *2.2. Climate parameters*

100 We obtained climate parameters from several databases based on the observation area  
101 latitudes and longitudes available in the dataset [1]. The data extraction and calculation of  
102 climate parameters were generally based on the procedures established in our previous  
103 publications [18, 20], which could be also accessed in our GitHub repository [25].

104 Climatic parameters with a spatial resolution of 2.5° were obtained from the WorldClim  
105 database (version 2.1) [26] for each observation area. In particular, we extracted the  
106 following monthly climate data according to the month of the median date in the data  
107 used for computing the growth rate: monthly mean temperature ( $T_{\text{mean}}$ ; °C), minimum  
108 temperature ( $T_{\text{min}}$ ; °C), maximum temperature ( $T_{\text{max}}$ ; °C), precipitation (mm), wind speed  
109 [ $\text{ms}^{-1}$ ], solar radiation (UV;  $\text{kJ m}^{-2}\text{day}^{-1}$ ), and water vapor pressure [kPa]. Moreover, we  
110 computed monthly diurnal temperature range (i.e.,  $T_{\text{max}} - T_{\text{min}}$ ; DTR; °C) and relative  
111 humidity based on  $T_{\text{mean}}$  and water vapor pressure. We also obtained the following annual  
112 climate parameters: temperature seasonality ( $T_{\text{seasonality}}$ ; standard deviation) and  
113 precipitation seasonality ( $P_{\text{seasonality}}$ ; coefficient of variation).

114 In order to evaluate the historical climate change, we computed warming velocity (WV)

115 [27, 28], defined as the temporal annual mean temperature (AMT) gradient divided by  
116 the spatial AMT gradient, where the temporal gradient is defined as the difference  
117 between the current and past AMT, available in the WorldClim database, and the spatial  
118 gradient was the local slope of the current climate surface at the observation area,  
119 calculated using the function *terrain* (with the option *neighbors = 4*) in the R package  
120 *raster* (version 2.9.5).

#### 121 2.4. Other related parameters

122 To investigate the effect of population density, we obtained 2020 population density (PD)  
123 data with a spatial resolution of 2.5' from the *Gridded Population of the World* (version  
124 4) [29].

125 To evaluate human impact, we used the human footprint (HF) scores, obtained from the  
126 *Last of the Wild Project* (version 3) [30]. The HF scores have a spatial resolution of 1 km  
127 grid cells and are defined based on human population density, human land use and  
128 infrastructure, and human access.

129 To evaluate the quality of human life, we used the gross domestic product (GDP) per  
130 capita and human development index (HDI), obtained from the *Gridded global datasets  
131 for Gross Domestic Product and Human Development Index over 1990-2015* [31]. HDI is  
132 defined based on life expectancy, education, and income (GDP per capita).

133 To evaluate the effect of travel restrictions, we manually extracted the dates when travel  
134 restrictions were imposed in each country from the Wikipedia page "*Travel restrictions  
135 related to the 2019–20 coronavirus pandemic*" [32]. The travel restrictions were classified  
136 into three categories: countries and territories implementing a global travel ban (Ban),  
137 countries implementing global quarantine measures (Qua), and non-global restrictions  
138 (NonG). When a country imposed multiple restriction types, the date when the strongest  
139 restriction was imposed was selected, where the order of the strength of travel restrictions  
140 was considered as follows: Ban > Qua > NonG. Many countries imposed travel  
141 restrictions after March 17, 2020 (see Figure S1 in our GitHub repository [25]). Thus, we  
142 considered a categorical variable (Ban) for the global travel restriction trend: 0 if the first  
143 date (see Section 2.1) is before March 17, 2020, and 1 otherwise.

#### 144 2.3. Data analyses

145 The statistical analyses were based on the procedures in [18, 20]. To evaluate the  
146 contribution of each variable to the growth rate, regression analysis was performed using  
147 R. Both ordinary least-squares (OLS) regression and the spatial analysis approach were  
148 considered. The dataset and R script for data analyses, used in this study, are available in  
149 our GitHub repository [25].

150 For the OLS regression, full models were constructed encompassing all explanatory  
151 variables ( $T_{\text{mean}}$ , DTR,  $T_{\text{seasonality}}$ , wind speed, precipitation,  $P_{\text{seasonality}}$ , UV, humidity, PD,  
152 HDI, WV, and Ban), and the best model was selected in the full model. The HF scores  
153 and GDP per capita were omitted because they were strongly correlated with PD and  
154 HDI, respectively. The best model was selected based on the sample-size-corrected

155 version of the Akaike information criterion (AICc) values using the R package *MuMIn*  
156 (version 1.43.6). Moreover, a model-averaging approach using *MuMIn* was adopted. The  
157 averaged model was obtained in the top 95% confidence set of models. A global Moran's  
158 test was performed to evaluate spatial autocorrelation in the regression residuals using the  
159 function *lm.morantest.exact* in the R package *spdep*, version 1.1.3.

160 PD and WV were log-transformed for normality. Precipitation and  $P_{\text{seasonality}}$  were square-  
161 root transformed for normality.  $T_{\text{mean}}$  was rescaled with  $\sqrt{(T_{\text{mean}} - 7.8)^2}$  to the  
162 quadratic relationship between temperature and transmission rate of COVID-19 [5, 7, 8,  
163 10, 11]. The quantitative variables were normalized to the same scale, with a mean of 0  
164 and a standard deviation of 1, using the *scale* function in R before the analysis.

165 For spatial analysis, a spatial eigenvector mapping (SEVM) modeling approach [21, 33]  
166 was also considered to remove spatial autocorrelation in the regression residuals.  
167 Specifically, the Moran eigenvector approach was adopted using the function  
168 *SpatialFiltering* in the R package *spatialreg* (version 1.1.5). As with the OLS regression  
169 analysis, full models were constructed, and then the best model was selected based on  
170 AICc values. The spatial filter was fixed in the model-selection procedures [33].

171 The contribution (i.e., non-zero estimate) of each explanatory variable to the growth rate  
172 of COVID-19 cases was considered significant when the associated  $p$ -value was less than  
173 0.05.

### 174 **3. Results and discussion**

175 The data in 300 areas were investigated (Figure 1). The OLS regression analysis (Table 1)  
176 and spatial analysis (Table 2) showed almost similar results because the statistical  
177 significances of spatial autocorrelations were moderate in the full model (Moran's  $I =$   
178  $0.077$ , and the associated  $p$ -value =  $0.021$ ) and best model ( $I = 0.084$ ,  $p = 0.027$ ) of the  
179 OLS regression analysis. The full, best, and averaged models showed almost similar  
180 results in both the OLS regression analysis and spatial analysis. The details of the results  
181 are as follows.

182 The temperature negatively correlated with the growth rate of COVID-19 cases. This  
183 indicates that high temperature (e.g., the arrival of summer season) reduces COVID-19  
184 transmission, consistent with several previous studies [4–11]. However, no humidity  
185 contribution was observed. This discrepancy might be due to differences in the datasets  
186 and data analyses between this study and previous studies. Previous studies (e.g., [4]),  
187 reported the association with humidity, was limited to the data on China; moreover, they  
188 used the measures based on the number of confirmed cases, although these measures may  
189 be affected by the difference of COVID-19 testing between areas. The contribution of  
190 humidity may be limited on a global scale. A similar tendency is observed in the case of  
191 influenza [13]; in particular, using specific humidity to determine transmission has a low  
192 predictive power at low- and mid-altitude sites, although humidity is believed to affect  
193 the transmission.

194 More importantly, however, we found that the growth rate was associated with the other  
195 parameters rather than temperature. In particular, we found that the growth rate of

196 COVID-19 cases showed a correlation with precipitation seasonality and warming  
197 velocity. Specifically, a lower growth rate was observed during a higher precipitation  
198 seasonality and a lower warming velocity; however, the contribution of precipitation  
199 seasonality was higher than that of warming velocity, according to the estimates of the  
200 models of the OLS regression analysis and spatial analysis. The observed associations  
201 may be reasonable in the context of the effects of seasonality and changing rapid weather  
202 variability on population dynamics of infectious diseases [17]. In particular, theory and  
203 experiment have indicated that climate seasonality can alter the spread and persistence of  
204 infectious diseases and that population-level responses can range from simple annual  
205 cycles to more complex multiyear variations. Therefore, climate seasonality and  
206 historical climate change can affect infectious disease transmission. In fact, rapid weather  
207 variability played a significant role in changing the strength of the influenza epidemic in  
208 the past [16]. However, the reason why temperature seasonality did not correlate with the  
209 growth rate remains unclear. Nevertheless, these results (the contribution of precipitation  
210 seasonality, in particular) may explain the exceptions (i.e., why the spreads of COVID-19  
211 are also observed in warm areas although previous studies suggest that high temperature  
212 reduces COVID-19 transmission). This may be because of the difference in precipitation  
213 seasonality between the observation areas. For example, the areas in Australia, Brazil,  
214 and Argentina were warm in March; however, they show low precipitation seasonality  
215 (Figure 2). Thus, the spreads might occur in these areas. Moreover, Europe and the USA  
216 might have undergone rapid spreads because they show low precipitation seasonality; on  
217 the other hand, the spread might have reached a peak relatively quickly in China because  
218 of relatively high precipitation seasonality.

219 The contribution of solar radiation is currently ambiguous. Solar radiation showed a  
220 positive association with the growth rate of COVID-19 cases. However, the results were  
221 less robust; in particular, the contribution was statistically significant in spatial analysis  
222 (Table 2), but not in the full and averaged models in the OLS regression (Table 1). Thus,  
223 it remains possible that the contributions partly observed in the analyses are artefacts.  
224 Assuming the positive association, the result is inconsistent with the fact that solar (UV)  
225 radiation is expected to reduce infection disease (e.g., influenza) transmission [13].  
226 Moreover, a pairwise correlation analysis showed no association between the growth rate  
227 and solar radiation (Spearman's rank correlation coefficient  $r = -0.06$ ,  $p = 0.31$ ).

228 The contributions of wind speed and precipitation were also limited. This is inconsistent  
229 with previous studies [8, 9]; however, statistical significances were not evaluated in these  
230 studies. This discrepancy might be due to differences in the data analyses between this  
231 study and previous studies. In particular, previous studies used the measures based on the  
232 number of confirmed cases; however, these measures may be affected by the difference  
233 of COVID-19 testing between areas. Hence, further examinations may be needed, given  
234 the importance of these climate parameters in infectious disease transmission [13, 17].

235 Non-climate parameters were also associated with the growth rate of COVID-19.  
236 According to the estimates of the models of the OLS regression analysis and spatial  
237 analysis, the contribution of travel restrictions was most significant than those of the  
238 climate parameters; in particular, travel restrictions showed a negative association with  
239 the growth rate. This result may be an extension of the result that the reduction of  
240 COVID-19 transmission due to interventions, including travel restrictions, in China [22,

241 23]. Our result implies that the travel restrictions in each country contributed to reducing  
242 COVID-19 transmission on a global scale.

243 The quality of human life (HDI) showed a positive association with the growth rate of  
244 COVID-19. This may be because HDI reflects life expectancy (i.e., areas with a higher  
245 HDI tend to have more older individuals because of a higher quality of human life).  
246 COVID-19 has the age specificity of cases and attack rates [34]; in particular, the  
247 epidemic risks of disease given exposure are likely to be the highest among adults aged  
248 from 50-69 years. Thus, the growth rate is expected to increase with HDI.

## 249 **4. Conclusions**

250 Intrigued by the question why COVID-19 transmission is observed in warm areas  
251 despites previous expectations of COVID-19 transmission reduction at high temperatures,  
252 we comprehensively investigated how several climate parameters are associated with the  
253 growth rate of COVID-19 cases and found that it was affected by precipitation  
254 seasonality and warming velocity rather than temperature. The effects were independent  
255 of population density, quality of human life, and travel restrictions. Our findings must  
256 necessarily be considered preliminary due to several limitations; in particular, it remains  
257 possible that the observed association is indirect. However, they may enhance our  
258 understanding of the COVID-19 transmission. As previous studies mentioned, high  
259 temperatures might reduce COVID-19 transmission. However, the effects may be  
260 restricted by intrinsic climate characteristics, such as precipitation seasonality and  
261 warming velocity. Moreover, the contributions of climate parameters to the growth rate of  
262 COVID-19 cases were moderate, while those of national emergency responses (i.e., travel  
263 restrictions) were more significant. Thus, slowing down the spread of COVID-19 due to  
264 the arrival of the summer season might not be expected. Instead, global collaborative  
265 interventions might be necessary to halt the epidemic outbreak.

## 266 **Ethics**

267 This study required no ethical permission.

## 268 **Data availability**

269 The datasets generated and analyzed in the current study are available in the GitHub  
270 repository: <https://github.com/kztakemoto/covid19climate>. The relevant R codes can be  
271 also found in the GitHub repository.

## 272 **Authors' contributions**

273 KT conceived and designed the study. KC and KT prepared the data and performed data  
274 analysis, interpreted the results, and wrote the manuscript. Both authors gave their final  
275 approval for publication.

## 276 **Competing interests**

277 There are no competing interests to declare.

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## Tables

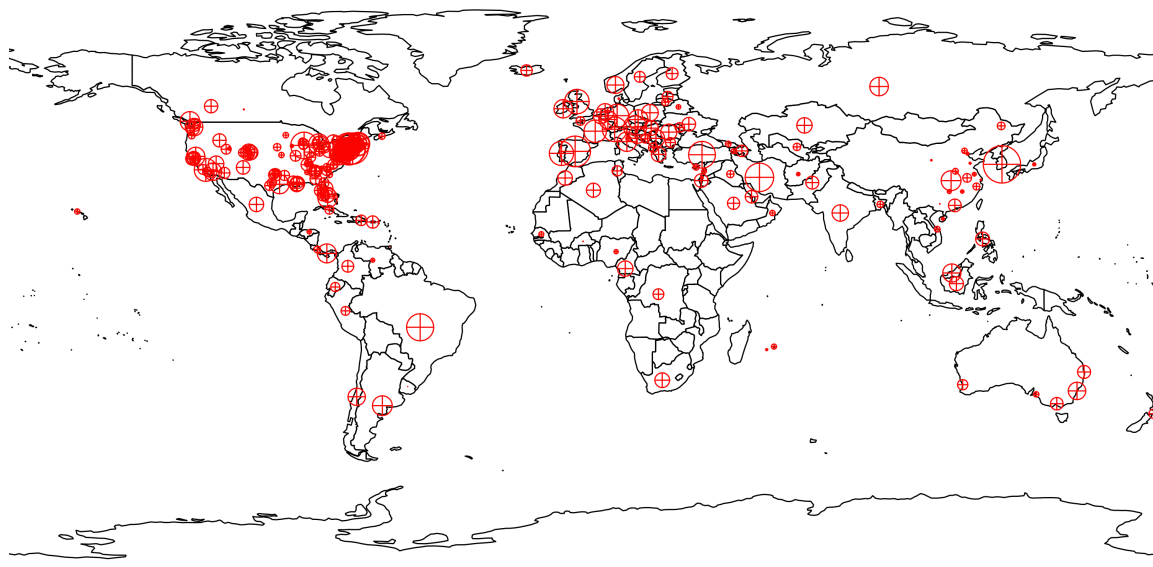
**Table 1. Influence of explanatory variables on the growth rate of COVID-19 cases based on the ordinary least squared regression approach.** The results of the full model, best model, and averaged model are shown, respectively. The abbreviations of variables are as follows:  $T_{\text{mean}}$  (monthly mean temperature), DTR (monthly diurnal temperature range),  $T_{\text{seasonality}}$  (temperature seasonality),  $P_{\text{seasonality}}$  (precipitation seasonality), UV (monthly solar radiation index), WV (warming velocity), PD (population density), HDI (human development index), and Ban (travel restrictions).  $R^2$  denotes the coefficient of determination for full and best models based on the OLS regression. SE is the standard error. Values in brackets are the associated  $p$ -values.

Variables	Full model			Best model			Averaged model		
	Estimate	SE	$p$ -value	Estimate	SE	$p$ -value	Estimate	SE	$p$ -value
$T_{\text{mean}}$	-0.18	0.07	0.014	-0.17	0.06	$9.0 \times 10^{-3}$	-0.16	0.07	0.032
Humidity	-0.10	0.09	0.27				-0.12	0.08	0.14
DTR	0.02	0.09	0.86				0.08	0.09	0.36
$T_{\text{seasonality}}$	-0.14	0.09	0.14				-0.13	0.09	0.15
Wind speed	-0.05	0.07	0.53				-0.04	0.07	0.57
Precipitation	-0.03	0.08	0.73				-0.01	0.08	0.90
$P_{\text{seasonality}}$	-0.30	0.10	$1.3 \times 10^{-4}$	-0.28	0.07	$7.2 \times 10^{-5}$	-0.30	0.08	$9.2 \times 10^{-5}$
UV	0.13	0.03	0.18	0.23	0.07	$5.4 \times 10^{-4}$	0.18	0.09	0.060
WV	0.18	0.07	$9.1 \times 10^{-3}$	0.14	0.06	0.017	0.15	0.07	0.028
PD	-0.06	0.06	0.27				-0.06	0.06	0.35
HDI	0.24	0.08	$9.1 \times 10^{-3}$	0.21	0.07	$1.8 \times 10^{-3}$	0.23	0.08	$1.7 \times 10^{-3}$
Ban	-0.68	0.13	$3.2 \times 10^{-7}$	-0.71	0.12	$1.8 \times 10^{-8}$	-0.68	0.13	$1.0 \times 10^{-7}$
Moran's $I$	0.077 (0.021)			0.084 (0.027)					
$R^2$	0.26 ( $2.1 \times 10^{-13}$ )			0.24 ( $1.2 \times 10^{-15}$ )					
AICc	791			783					

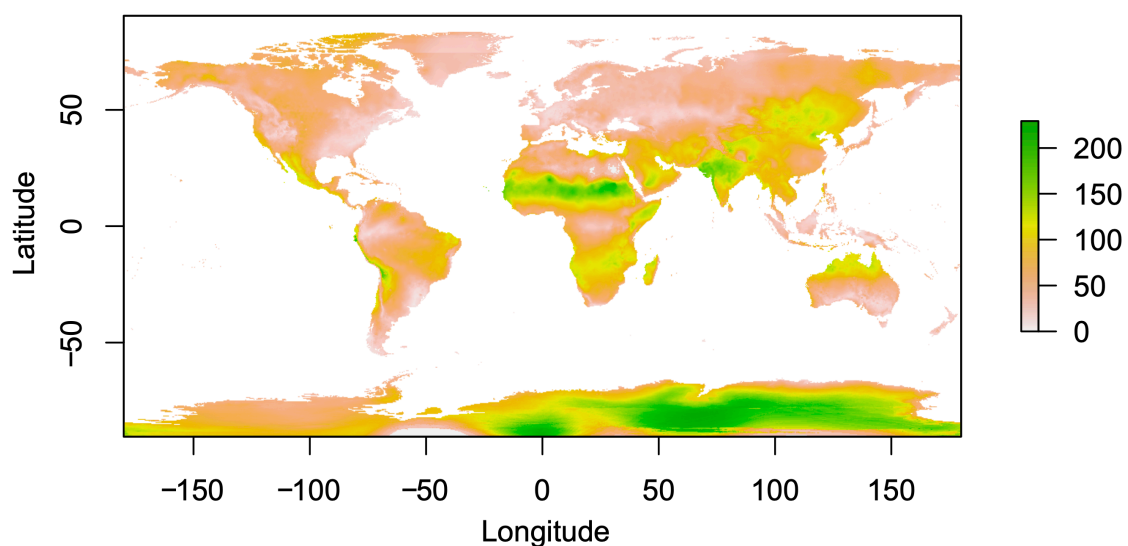
**Table 2. Influence of explanatory variables on the growth rate of COVID-19 cases based on the spatial analysis approach.** The results of the full model, best model, and averaged model are shown.  $R^2$  denotes the coefficient of determination for full and best models based on the SEVM modelling. SE is the standard error. Values in brackets are the associated  $p$ -values. See Table 1 for description of table elements.

Variables	Full model			Best model			Averaged model		
	Estimate	SE	$p$ -value	Estimate	SE	$p$ -value	Estimate	SE	$p$ -value
$T_{\text{mean}}$	-0.20	0.07	$4.8 \times 10^{-3}$	-0.21	0.06	$1.4 \times 10^{-3}$	-0.20	0.07	$3.1 \times 10^{-3}$
Humidity	-0.03	0.09	0.73				-0.05	0.08	0.55
DTR	0.07	0.09	0.47				0.09	0.07	0.21
$T_{\text{seasonality}}$	-0.01	0.09	0.89				0.00	0.09	0.99
Wind speed	-0.03	0.07	0.68				-0.04	0.07	0.59
Precipitation	0.02	0.08	0.77				0.01	0.08	0.87
$P_{\text{seasonality}}$	-0.32	0.08	$1.2 \times 10^{-4}$	-0.31	0.07	$1.5 \times 10^{-5}$	-0.33	0.08	$2.1 \times 10^{-5}$
UV	0.24	0.10	0.013	0.30	0.07	$2.1 \times 10^{-5}$	0.27	0.08	$1.0 \times 10^{-3}$
WV	0.19	0.07	$3.7 \times 10^{-3}$	0.20	0.06	$8.7 \times 10^{-4}$	0.19	0.06	$2.8 \times 10^{-3}$
PD	-0.03	0.06	0.55				-0.04	0.06	0.49
HDI	0.21	0.08	$9.3 \times 10^{-3}$	0.21	0.07	$2.1 \times 10^{-3}$	0.21	0.07	$3.1 \times 10^{-3}$
Ban	-0.73	0.13	$2.5 \times 10^{-8}$	-0.72	0.12	$4.8 \times 10^{-9}$	-0.73	0.12	$< 2.0 \times 10^{-16}$
Moran's $I$	-0.052 (0.51)			-0.054 (0.60)					
$R^2$	0.33 ( $2.9 \times 10^{-16}$ )			0.32 ( $< 2.2 \times 10^{-16}$ )					
AICc	773			763					

## Figures



**Figure 1. Distribution of the observation areas included in this study.** Red symbols indicate the observation areas. Symbol size indicates the growth rate of COVID-19 cases.



**Figure 2. World distribution of precipitation seasonality.**